Accepted Manuscript

Global reconstructed daily surge levels from the 20th Century Reanalysis (1871–2010)

Alba Cid, Paula Camus, Sonia Castanedo, Fernando J. Méndez, Raúl Medina

 PII:
 S0921-8181(16)30494-5

 DOI:
 doi: 10.1016/j.gloplacha.2016.11.006

 Reference:
 GLOBAL 2513

To appear in: Global and Planetary Change

Received date:10 August 2015Revised date:13 May 2016Accepted date:9 November 2016

Please cite this article as: Cid, Alba, Camus, Paula, Castanedo, Sonia, Méndez, Fernando J., Medina, Raúl, Global reconstructed daily surge levels from the 20th Century Reanalysis (1871–2010), *Global and Planetary Change* (2016), doi: 10.1016/j.gloplacha.2016.11.006

This is a PDF file of an unedited manuscript that has been accepted for publication. As a service to our customers we are providing this early version of the manuscript. The manuscript will undergo copyediting, typesetting, and review of the resulting proof before it is published in its final form. Please note that during the production process errors may be discovered which could affect the content, and all legal disclaimers that apply to the journal pertain.



Global reconstructed daily surge levels from the 20th Century Reanalysis (1871-2010)

Alba Cid^{a,*}, Paula Camus^a, Sonia Castanedo^b, Fernando J. Méndez^b, Raúl Medina^a

^a Environmental Hydraulics Institute "IH Cantabria", Universidad de Cantabria. Spain ^b Departamento de Ciencias y Técnicas del agua y del Medio Ambiente, Universidad de Cantabria. Spain

Abstract

Studying the effect of global patterns of wind and pressure gradients on the sea level variation (storm surge) is a key issue in understanding the recent climate change effect on the dynamical state of the ocean.

The analysis of the spatial and temporal variability of storm surges from observations is a difficult task to accomplish since observations are not homogeneous in time, scarce in space, and moreover, their temporal coverage is limited. A recent global surge database developed by AVISO (DAC, Dynamic Atmospheric Correction) fulfilled the lack of data in terms of spatial coverage, but not regarding time extent, since it only includes the last two decades (1992-2014).

In this work, we use the 20th Century reanalysis V2 (20CR), which spans the years 1871 to 2010, to statistically reconstruct daily maximum surge levels at a global scale. A multivariate linear regression model is fitted between daily mean ERA-interim sea level pressure fields and daily maximum surge levels from DAC. Following, the statistical model is used to reconstruct daily surges using mean sea level pressure fields from 20CR. The verification of the statistical model shows good agreements between DAC levels and the reconstructed

^{*}Corresponding author at: Environmental Hydraulics Institute, "IH Cantabria", Universidad de Cantabria, C/Isabel Torres n 15, Parque Científico and Tecnológico de Cantabria, 39011, Santander, Spain. Tel.:+34 942 201414

Email address: alba.cid@unican.es (Alba Cid)

surge levels from the 20CR. The validation of the reconstructed surge with tide gauges, distributed throughout the domain, shows good accuracy both in terms of high correlations and small errors. A time series comparison is also depicted at specific tide gauges for the beginning of the 20th century, showing a high concordance.

Therefore, this work provides to the scientific community, a daily database of maximum surge levels; which correspond to an extension of the DAC database, from 1871 to 2010. This database can be used to improve the knowledge on historical storm surge conditions, allowing the study of their temporal and spatial variability.

Keywords:

Climate variability, multiple linear regression, statistical modelling, storm surge, historical reconstruction

1. Introduction

The storm surge is one of the main variables that describe the dynamical state of the ocean. It is defined as the sea level variation due to wind stress and
sea level pressure gradients over the sea surface. The storm surge magnitude has a large spatial variability and can reach extremely high values associated to
tropical and extra-tropical storms (extreme storm surge events).

The storm surge is one of the sea level components, and can be extracted from tide gauge records; but since many of the world areas are ungauged, the global coverage is scarce. Besides, areas where instrumental records are available, normally present short time records and hence only recent analysis can be carried out. Nevertheless, there is a global assessment from long centennial records (Marcos et al., 2015) focus on sea level extremes, although it is limited to areas with available tide gauges. One option to overcome these shortcomings is the use of numerical models. They enable us to simulate the magnitude

- of the storm surges at a global scale or in an specific area. In fact, there
- ¹⁶ is a global surge database, the Dynamic Atmospheric Correction from AVISO

(hereinafter DAC database, http://www.aviso.altimetry.fr/), that provides
global records from 1992 and onwards (a description of the numerical configuration can be found in Carrère and Lyard 2003). Nevertheless, DAC database is
not long enough to address inter-decadal climate variability or likely long-term trends. The numerical modelling at a global scale and for a long period of time

²² requires an extremely high computational effort.

- An alternative to numerical models is the statistical modelling; different techniques of statistical reconstruction are widely used in climate or ocean variables. Statistical downscaling techniques can be classified into: transfer func-
- tions, weather-type approaches and stochastic weather generators (Giorgi et al., 2001). Regarding marine climate, linear transfer functions (regression models)
- have been applied to downscale the significant wave height at global (Wang et al., 2012) and at regional scale (Casas-Prat et al., 2014). Also, a weather-
- ³⁰ type model has been proposed to downscale multivariate wave climate (Camus et al. 2014b, Espejo et al. 2014).
- The application of statistical downscaling approaches is usually limited to specific locations. Calafat and Gomis (2009) used a reduced space optimal
- ³⁴ interpolation analysis to reconstruct the sea level in the Mediterranean Sea. For storm surges specifically, Dangendorf et al. (2014) compared, in the North
- ³⁶ Sea, the long-term behaviour of surges to that of reanalysis wind fields by means of a statistical-empirical formulation (Müller-Navarra and Giese, 1999).
- The goal of this study is to perform a global reconstruction of the storm surge by means of a statistical model in order to extend the DAC temporal coverage.
- ⁴⁰ For achieving this purpose, we define the statistical relationship between the storm surges from DAC database and their drivers (pressure and wind fields
- ⁴² from ERA-interim reanalysis, Dee et al. 2011) using multiple linear regression. In this work, although using the term storm surge, we are not reconstructing
- ⁴⁴ specific strong events but to a continuous time series of daily values. Once the statistical model is calibrated and verified, we use a global atmospheric database
- that starts at the end of the 19th century (20th century reanalysis V2, Compo et al. 2011) to reconstruct the storm surge at a global scale and for a long period

⁴⁸ of time (1871-2010).

The work is structured as follows. Section 2 describes the three global databases and the tide gauges used in this study. The methodology of the statistical model is explained in depth in section 3 and its accuracy is shown in section 4, where the verification results are depicted. Finally, the global reconstruction of the surges and its validation with tide gauges is shown in section 5.

54 Main conclusions are summarised in section 6.

2. Databases description

- The surge database corresponds to the Dynamic Atmospheric Correction (DAC), produced by CLS Space Oceanography Division using the MOG2D model from Legos and distributed by Aviso, with support from Cnes (http: //www.aviso.altimetry.fr/). MOG2D (2 Dimensions Gravity Waves model)
- ⁶⁰ is a finite element, barotropic, non-linear, two-dimensional shallow water hydrodynamic model, derived from Lynch and Gray (1979). The model is forced by
- ⁶² pressure and wind fields from the European Centre for Medium-range Weather Forecasts (ECMWF) analysis, with a temporal resolution of 6 hours and includ-
- ⁶⁴ ing shallow water areas and marginal seas. Barotropic sea level outputs span from September 1992 to present and are provided on a regular grid of 0.25°x0.25°
- every 6 h. The operational DAC database is made of the high frequencies (i.e less than 20 days) obtained from MOG2D barotropic model and the low fre-
- quencies of the inverse barometer (IB) assuming a static response of the ocean to the atmospheric forcing, and neglecting wind effects for low frequency (i.e
- ⁷⁰ more than 20 days). Therefore, a 20 days high pass (respectively low pass) filtering is applied to separate high and low frequencies (see Eq. 1).

$$DAC = MOG2D(T < 20days) + IB(T > 20days)$$
(1)

Regarding the atmospheric forcing, sea level pressure fields (SLP) were selected from two different global atmospheric databases: ERA-Interim reanalysis

- 74 (Dee et al., 2011) from the ECMWF, which is the DAC forcing field, and the twentieth Century Reanalysis V2 (20CR, Compo et al. 2011).
- ⁷⁶ SLP fields from ERA-Interim, consist of 6-hourly atmospheric data at 0.75° of spatial resolution and spanning from 1 January 1979 to present. In this work,
- only data covering DAC period (1992 2014) are selected. These data are used to calibrate and verify the statistical model (see Section 3).
- SLP fields from the 20CR are available every 6 hours at a spatial resolution of 2°, covering the period between 1871 and 2010 (i.e. 140 years). These data
 are used for the statistical reconstruction of surge levels (see Section 5).
- Concerning instrumental data, all available tide gauges from the University
- of Hawaii Sea Level Center (UHSLC, http://uhslc.soest.hawaii.edu/data/ download/rq) were downloaded at hourly scale. From the total amount of
- 643 tide gauges, only those with more than one year of data before 2010 are used to validate the reconstructed surge from 20CR. In order to compare both
- signals, the tide gauge residuals are obtained by subtracting the astronomical tide (computed using t-tide, Pawlowicz et al. 2002) to the hourly values, and
- ⁹⁰ subtracting a 30-day moving average to both, modelled and measured data. This leads to a validation of the storm surge reconstruction at 386 tide gauge
- 92 locations distributed worldwide.

3. Statistical modelling methodology

94 3.1. Predictor and predictand definitions

The aim of the statistical reconstruction is to estimate surge levels (predictand) from local atmospheric conditions (predictor) based on a statistical relationship. Specifically, our interest consists in finding the statistical rela-

- ⁹⁸ tionship between mean daily atmospheric conditions and maximum daily surge levels. Following this purpose, the predictand is defined as the maximum of
- the 4 daily DAC values (DAC has a 6-hourly temporal resolution) at each grid point. The spatial resolution of the statistical reconstruction is determined by
- ¹⁰² the 2° resolution of the 20CR. This spatial resolution is considered sufficient

to represent the surge variability at a global scale. Therefore, surge levels are ¹⁰⁴ selected from DAC database every 2°.

As mentioned in Section 2, SLP fields from ERA-interim are used to obtain the statistical relationship between these atmospheric fields and surge levels. Although both, surface wind and pressure fields, are the drivers of the storm surge, in this work we are only extracting SLP fields from the atmospheric reanalysis. This is because in global circulation models, sea wind fields are not as well reproduced as sea level pressure fields (Wang et al., 2010); but since the geostrophic wind speed is proportional to the square pressure gradient, SLP gradients are calculated, and taken into account in the statistical modelling, to also have an estimation of the wind speed.

Although SLP from ERA-interim reanalysis has a horizontal resolution of 0.75°, SLP fields and the calculated gradients have been re-arranged in a 2°

¹¹⁶ grid at a daily scale. The final step in the definition of the predictor is the establishment of the spatial coverage. Taking each of the predictand grid points,

a local area of 4°x4° enclosing this target point is defined. As a result, the predictor is composed of 9 SLP values and 9 SLP gradients (18 components)
centred in the predictand location.

The same process is carried out for the definition of the predictor from 20CR ¹²² SLP fields, with the only difference that in this case, data are already in a 2° grid.

Therefore, the predictand consists of daily maximum values of surge levels from September 1992 to December 2014 in a 2° global grid; the predictor for

the statistical model configuration consists of daily means of SLP fields and square SLP gradients from ERA-interim, covering an area of $4^{\circ}x4^{\circ}$ centred at the

predictand location. Once the statistical model is defined, 20CR daily predictors are used to reconstruct surge levels from 1871 to 2010 at each location (10594
locations).

3.2. Description of the statistical model

The statistical method chosen to reconstruct surge levels is based on the method used in Camus et al. (2014a) to downscale multivariate wave climate.
It comprises a multivariate regression model fitted between daily maximum surge level (predictand) and the principal components(PCs) of the mean daily
SLP and gradients (predictor).

The first step in the methodology consists in performing a principal compo-¹³⁸ nent analysis (PCA) of the predictor to reduce the dimensionality of the problem while preserving the maximum variance of the data sample. PCA projects the ¹⁴⁰ original data on a new space, searching for the maximum variance of the sam-

ple data. The eigenvectors (empirical orthogonal functions, EOFs) of the data

- ¹⁴² covariance matrix define the vectors of the new space. The transformed components of the original data over the new vectors are the principal components
- (PCs). The original predictor, which varies with space and time X(x, t), can be expressed as a linear combination of EOFs (accounting for the space variabil-
- ity) and PCs (accounting for the time variability). Eq.2 shows that at any given time, the predictor can be estimated as the spatial pattern detected (EOF) mul-
- tiplied by the corresponding coefficient for that instant (PC), adding all the N components (dimensions).

$$X(x,t_i) = EOF_1(x) \times PC_1(t_i) + EOF_2(x) \times PC_2(t_i) + \dots + EOF_N(x) \times PC_N(t_i)$$
(2)

where N is the number of dimensions in the data, specifically in this case N = 18.

Prior to PCA, SLP and SLP gradients are standardise to avoid attributes
¹⁵² with large values (and larger variances) to dominate the analysis. Then, PCs of the predictor have been calculated for the period 1992 - 2014. Following, we
¹⁵⁴ have selected the components that explain the 95% of the variance. The number of PCs explaining the 95% of the variance varies spatially between 4 and 12 (not

¹⁵⁶ shown) from the total amount of 18 components. Therefore, although in this case the dimension reduction is not highly relevant, PCA is carried out anyway

- ¹⁵⁸ so that this methodology could also be applied to wider predictors (i.e higher dimensions).
- The next step consists of establishing the relationship between the predictor PCs that explain the 95% of the variance and the surge levels (predictand). The multivariate regression model is fitted in a forward procedure: first predictor PC is obtained from the best fit (smallest sum of squared errors, SSE) among each of the components separately ($PC_{(1)}$ in Eq.3). The second predictor PC ($PC_{(2)}$ in Eq.3) is chosen from the rest of the PCs so that it gives the best
- fit with two predictors, the best predictor selected in the previous model plus one of the remaining potential predictors. The cycle continues until a more
- $_{168}$ complex model does not produce a significant improvement (at the 5% level of significance) in the multivariate regression fit. This evaluation is based on the F
- ¹⁷⁰ statistics that compare the SSE of fitting a simpler-parameter model with that of a more complex parameter model (Wang et al., 2010). Therefore, although all
- PCs explaining the 95% of the variability could be used in the regression model, only those that produce a significant improvement are taken into account.
- ¹⁷⁴ A leave-one-out-cross-validation process is used to set up the statistical model: The regression model is fitted for all years except one, which is used
- to validate the surge reconstruction. Then, the regression model is fitted 22 times (there are 22 complete years of surge levels, 1993-2014). Results of the
 model verification are shown in Section 4.

This methodology allows us to estimate surge levels as a linear combination of the most important PCs. As can be seen from Eq.3, surge levels at any given location (x_i) can be estimated from a specific number of PCs, which varies throughout the domain.

$$surge(x_i, t) = a_i + b_{1,i} \times PC_{(1)}(x_i, t) + b_{2,i} \times PC_{(2)}(x_i, t) + \dots + b_{n,i} \times PC_{(n)}(x_i, t)$$
(3)

where n is the number of PCs that achieved a statistical improvement of the results (following F statistics) and selected in a forward procedure from the

PCs that explained the 95% of the variance. $a_i, b_{1,i}, ..., b_{n,i}$ are the coefficients obtained in the regression model.

Once the statistical model is defined, SLP and gradients from the 20CR reanalysis are standardised and projected into the EOFs detected for ERAinterim to obtain the PCs for the 20CR predictors. Finally, using the coefficients

¹⁹⁰ and the PCs identified as in Eq.3, daily surge values can be reconstructed for the 1871-2010 period.

¹⁹² 4. Model verification

The verification of the statistical model shows the quality of the fittings in terms of correlation coefficient (Pearson coefficient, ρ), root mean square error 194 (RMSE), and the RMSE relative to the maximum surge variability at each location $(RMSE_{relative})$. Fig.1 shows the spatial distribution of the fitting 196 quality. As can be seen from Fig.1a, correlation coefficients are above 0.65 all along the domain except for the areas located along the north-atlantic coast 198 of South America, the area from the Gulf of Thailand through Indonesia to North Australia, and along the Siberian and Beaufort Seas. An extended area 200 of relative low correlations (around 0.7) is located along tropical areas. The explanation for this lies in the fact that the ocean dynamics at high latitudes 202 have smaller space and time scales. The predictor used in the proposed statistical model describes better the dynamical response of the storm surge in those 204 areas in comparison with tropical areas. Besides, the small magnitude of the storm surge in tropical areas is more difficult to reproduce. Fig.1b shows that 206 the RMSE is generally higher in the Northern Hemisphere (maximum values of ~ 10 cm), specifically in semi-enclosed areas as the Hudson bay or the North 208 Sea, as well as along high latitudes of Russian and Siberian coasts. This is because the storm surge generation at shallow waters, especially in semi-enclosed 210

effect of the bathymetry is taken into account by the statistical model in the predictand term (storm surge), although the non-linear interactions of the storm

areas, is highly related to local wind conditions and the local bathymetry. The

- ²¹⁴ surge with the bathymetry can be partially misrepresented. Therefore, a higher resolution predictor should be required to improve the skill of the statistical
- ²¹⁶ model at those areas. Nevertheless, a generalised predictor is used in this study due to the global scale of the storm surge reconstruction. Fig.1c displays the
- 218 RMSE relative to the maximum surge variability detected at each grid point. This reveals areas of maximum relative errors of around 10% along equatorial
- regions, not detected in Fig.1b due to the small magnitude of the surge levels along the tropical areas (see Fig.4).
- A comparison, between the reconstructed surge and DAC, at a few specific points is also carried out (red dots in Fig.2). Fig.3 shows the time series compar-
- ²²⁴ ison and the scatter plots at these 6 locations, contrasting DAC data (red lines) with the statistical reconstruction using ERA-interim predictors (blue lines). It
- can be seen how the equatorial area (point 4) presents the lowest correlations and the highest relative errors, even so, ρ reaches a 0.86 value and the relative
- error is around 6%. A general good fitting to the bisector can be seen at all grid points. Fig.3 confirms that the poorest agreement is located at equatorial
- ²³⁰ areas, but although absolute values are not exactly reproduced at grid point 4, mean levels are well represented. It is also important to notice that the surge
- magnitude at these tropical zones is about one order of magnitude lower than at the points located in other areas. Time series at the rest of grid points follow
- ²³⁴ DAC data properly.

Last comparison between DAC data and the statistical reconstruction using
ERA-interim predictors is represented in Fig.4. It shows the spatial value of the 99.5% percentile obtained from the complete daily surge series. As can be
seen, the spatial pattern is almost identical; main differences are found along the North European coastline, where the statistical reconstruction does not
reach the values displayed in the original data. As we mentioned previously, the generation of the storm surge at coastal areas is mainly forced by local

- 242 winds, for which the resolution of the predictor is not sufficient to resolve local features. Maximum surge values are present in these areas, reaching magnitudes
- 244 of about 1 m. As already remarked, smallest surges are located around the



Figure 1: Spatial verification of the statistical reconstruction. a) correlation coefficient(ρ). b) RMSE (cm). c) RMSE relative to the maximum surge variability(%).



Figure 2: Selected points where a point comparison is carried out. Red grid points are used to validate the statistical model and to compare with the 1871-2010 reconstruction. Blue grid points are compared with tide gauges. Black dots show the spatial resolution of 2°.

equatorial area, where 99.5% is even negative due to the fact that these areas have high mean pressure conditions (permanent anticyclonic conditions). The extratropical Southern Ocean (40°S - 60°S) is an area with relatively higher

- values, where the 99.5% percentile is above 60 cm. Wind conditions at high latitudes of the Southern Hemisphere are the most energetic at a global scale;
- 250 the general conditions of low pressures present in this area also contribute to maintain high surge levels.



Figure 3: Surge (m) time series comparison and scatter plots at the 6 grid points defined in Fig.2. Red line represents DAC data, blue line corresponds to the statistical reconstruction using ERA-interim predictors. (1): $\rho = 0.98 RMSE = 1.7 \text{ cm}$; (2): $\rho = 0.93 RMSE = 2.7 \text{ cm}$; (3): $\rho = 0.96 RMSE = 1.4 \text{ cm}$; (4): $\rho = 0.86 RMSE = 1 \text{ cm}$; (5): $\rho = 0.95 RMSE = 3.6 \text{ cm}$; (6): $\rho = 0.93 RMSE = 1.2 \text{ cm}$



Figure 4: Value (m) of the 99.5% percentile for DAC data (top) and for surge reconstruction using ERA-interim predictors (bottom).

²⁵² 5. Global reconstruction of surges (1871-2010)

In this section, a global statistical surge reconstruction is performed using 254 20CR predictors, obtaining maximum daily surge levels for the period 1871-2010.

²⁵⁶ 5.1. Comparison during the control period (1992-2010)

A comparison between DAC data and the reconstructed 20CR surge is carried out for a control period (1992-2010) where data from both sources coincide. 258 Fig.5 shows the spatial comparison, same as in Fig.1 but using 20CR predictors instead of those from ERA-interim. The spatial pattern of both figures is very 260 similar for the 3 statistical indicators. Main differences between Fig.1 and Fig.5 are located in the Southern Hemisphere, specifically in the Pacific Ocean close 262 to Antarctica, where the correlation coefficient decreases from 0.95 (Fig.1a) to 0.85 (Fig.5a) and the relative error is doubled (even so, barely an 8%). These 264 differences with the statistical verification can only be due to the quality of the 20CR reanalysis. The higher quality of the 20CR upper-air fields over the 266 whole Northern Hemisphere and the mid-latitudes of the Southern Hemisphere has been highlighted when comparing with ERA-40 reanalysis (Compo et al., 268 2011). Fig.6 shows the time series comparison between DAC data (red lines) and the statistical reconstruction using 20CR predictors (black lines) for the 6 270 grid points numbered in Fig.2. Grid points 4 (equatorial area) and 6 (Indian Ocean) show extreme storm surge events in the reconstruction that are not 272 present in the original data. This can also be seen from the scatter plots; for those points, an overestimation of the reconstructed surge values regarding DAC 274 data is noticed. The rest of the points show a good agreement, with correlation coefficients and errors of about the same magnitude as in Fig.3. 276 Comparing SLP data from 20CR and ERA-interim (not shown) we verified

that the aforementioned extreme surges are due to extreme falls in the SLP of 20CR reanalysis that are not consistent with the SLP in ERA-interim. A
brief study has shown us that those drops in the SLP correspond to tropical

cyclones (TC), included in 20CR but not in ERA-interim. In the 20CR, the
estimated minimum central pressure observations for TCs from the International
Best Track Archive for Climate Stewardship (IBTrACS) are assimilated into
the reanalysis model. This assimilation may justify the more intense minimum
pressures found in 20CR, when comparing to ERA-Interim.

²⁸⁶ 5.2. 99.5% percentile of daily surge levels

The value of the 99.5% percentile is calculated for the statistical reconstruction using 20CR predictors (see Fig.7). It is done for two periods, one which 288 corresponds to the control period (1992-2010) and that will allow us to compare the results with the DAC data and the other one which covers the complete 290 20CR period (1871-2010). As can be seen from Fig.7, the 99.5% spatial pattern for the control period (top panel) is very similar to the one corresponding to 202 DAC (top panel in Fig.4). When taking into account the complete period, surge levels decrease generally throughout the domain. The decrease is obvious in the 294 Antarctic surrounding area, where the 99.5% value decreases from around 80 cm to 60 cm. A decrease of 10 cm is found around Canada Arctic coasts. The 296 lower magnitude of the storm surge 99.5% percentile, for the period 1871-2010, is in agreement with positive trends in the extra-tropical cyclone activity in 298 the Northern Hemisphere and significantly in the Southern Hemisphere (Wang et al., 2013). However, long-term trends in northeast Atlantic storminess de-300 rived from 20CR and observations are found to be inconsistent in the first half of the twentieth century due to inhomogeneities in 20CR, mostly caused by lower 302 data assimilation (Krueger et al., 2013). Inconsistencies between storm surge

³⁰⁴ record and storm surge reconstruction using 20CR are detected only before 1910 in the North Sea (Dangendorf et al., 2014).

306 5.3. Validation with tide gauges

The surge reconstruction from 20CR predictors has also been compared to measures from tide gauges. The closest grid point to each tide gauge has been



Figure 5: Spatial comparison during the control period between DAC levels and the reconstructed surge using 20CR predictors. a) correlation $\operatorname{coefficient}(\rho)$. b) RMSE (cm). c) RMSE relative to the maximum surge variability(%).



Figure 6: Surge (m) time series comparison at the 6 grid points defined in Fig.2. Red line represents DAC data, black line depicts the surge reconstruction using 20CR predictors. (1): $\rho = 0.97 \ RMSE = 1.8 \text{ cm};$ (2): $\rho = 0.92 \ RMSE = 2.8 \text{ cm};$ (3): $\rho = 0.95 \ RMSE = 1.7 \text{ cm};$ (4): $\rho = 0.71 \ RMSE = 1.4 \text{ cm};$ (5): $\rho = 0.93 \ RMSE = 4.3 \text{ cm};$ (6): $\rho = 0.9 \ RMSE = 1.4 \text{ cm}$



Figure 7: 99.5% percentile value (m) of the statistical reconstruction using 20CR for the period 1992-2010 (top) and for the complete 20CR period 1871-2010 (bottom)

K

	selected from the reconstructed surge database. Fig. 8 presents the valida-
310	tion results for all available tide gauges, where the size of each coloured dot
	depends on the length of the tide gauge record; small dots, as for example
312	Ammassalik in Greenland, correspond to tide gauges with less than 5 years of
	measures; medium size dots, as for instance Reykjavik in Iceland, correspond
314	to tide gauges with more than 5 but less than 50 years of records; large dots,
	as Lerwick in Scotland, are tide gauges with more than 50 years of measures.
316	The correlation coefficient depicted in Fig. 8a shows that minimum values are
	above 0.5 except for the tide gauges located in tropical zones. The lowest val-
318	ues are found in Central America, the north of South America and the Gulf
	of Thailand. Higher correlations are found in extra-tropical areas with values
320	that are generally above 0.7. RMSE (see Fig. 8b) are low at equatorial areas,
	due to small magnitude of the surge, and up to 15 cm at the south-east coasts
322	of South America and semi-enclosed areas of the North Sea. Relative errors
	shown in Fig. 8c reach higher values (up to 20 $\%)$ in islands mainly spread
324	through tropical areas, while at the coasts of the Northern Hemisphere, relative
	errors are below 5 %. This spatial pattern follows the previously found during
326	the verification process, strengthening the fact that the storm surge in tropical
	zones is more difficult to reproduce. It is worth noting that these differences
328	can also be due to the different dynamics that are gathered in each surge series.
	The reconstructed signal represents the sea level variation due to meteorolog-
330	ical factors only, while the daily measures from tide gauges also account for
	signals due to oceanographic processes (non-tidal residual). The importance of
332	the oceanographic processes is spatially variable throughout the global ocean
	(Woodworth and Menéndez, 2015).
334	A time series comparison is also shown (see Fig.9) for six of the longest
	UHSLC records (blue dots in Fig.2. i: Cuxhaven, ii: Brest, iii: Newlyn, iv:
336	San Francisco, v: Honolulu, vi: Fort Denison, Sidney), where daily maximum
	of the tide gauges (red line) are compared to the reconstructed daily maximums

from 20CR (black line) for the end of the nineteen century or the beginning of the twenty century. Time series from Brest and Newlyn coincide extremely well,

both in the timing and the magnitude, with a correlation coefficient of 0.89, and 340 a RMSE of 6 cm; Cuxhaven, San Francisco and Fort Denison show a slightly lower agreement (ρ is 0.8, 0.75 and 0.6, respectively. RMSE is about 25, 5 and 342 7 cm, respectively); Honolulu shows periods where the concordance is low (ρ is 0.41 and RMSE is 3 cm). Assuming that DAC is of high quality all over the 344 domain, this is the expected result since Brest and Newlyn are located in areas with remarkably high levels of agreement (see Fig.5); Cuxhaven in the North 346 Sea and San Francisco, are placed in areas with slightly smaller correlation coefficients and higher relative errors; Honolulu is located in an area where the 348 lower agreements are found. Besides, DAC performances are also better for coastal tide gauges rather than islands tide gauges (Carrère and Lyard, 2003). 350

From the comparison with these six tide gauges, it can be seen that the reconstructed storm surge database is able to represent the measured surge

signal, also from periods back at the end of the nineteen century with relatively

354 high accuracy.



Figure 8: Validation of the reconstructed surge with tide gauges. a) correlation coefficient(ρ). b) RMSE (cm). c) RMSE relative to the maximum surge variability(%).



Figure 9: Time series comparison of the reconstructed daily maximum surges (black line) with daily values from tide gauges (red line) at 5 locations (blue dots in Fig.2).

6. Conclusions

This study shows a novel combination of the use of numerical and statistical 356 models, providing a global storm surge database that estimates daily maximum surge levels from 1871 to 2010, at a spatial resolution of 2°. It has been carried 358 out using a statistical model, based on multiple linear regression, that relates the mean SLP and gradients from ERA-interim reanalysis with daily maximum 360 surge levels from DAC database. The model is calibrated and validated, and then used to reconstruct daily maximum surge levels, taking the atmospheric 362 SLP fields from the 20CR as predictors. The verification of the model showed an extremely high performance all 364 over the world, with minimum agreements (maximum $RMSE_{relative}$ of 10% and minimum ρ of 0.65) around equatorial zones and semi-enclosed areas. Two 366 different reasons explain these lower performance areas. In the case of semienclosed areas, a finer predictor resolution would be needed to capture the 368 variability of the atmospheric structures and hence to properly reproduce the surge; in equatorial areas, the small surge magnitude and also the large temporal 370 and spatial scales of the ocean dynamics make more difficult for the statistical model to accurately reproduce the surge. 372 The comparison between the 20CR reconstructed surge and the original data for the control period, shows a decrease in the agreement at the Southern 374 Hemisphere, specially around the Antarctic area. This is explained in terms of the lower quality of the 20 Century Reanalysis in the Southern Hemisphere 376 (Compo et al., 2011). Around tropical zones, we find extreme data in the

- 378 reconstruction using 20CR, that correspond to tropical cyclones and that are not present in the DAC database. This is due to the fact that the 20CR assimilates
- information about tropical cyclone tracks that are not gathered in the ERA-interim reanalysis, since the detected minimum pressures are more noticeable
 in 20CR than in ERA-interim.

The validation with a large number of tide gauges has shown the variable ³⁸⁴ spatial quality of the reconstructed surge, and the time series comparison with

specific tide gauges has also proven the ability of the model to reproduce the
³³⁶ surge signal far from the control period. Nevertheless, the extreme values are often underestimated. This because, although not shown, the comparison of
the tide gauge residuals with DAC, gives us a very similar picture to the one presented in Fig.8, meaning that the selected predictand database does not
reproduce maximum values accurately.

It is important to note, that at high latitudes there is a variable ice coverage that would eliminate the surge signal, but since DAC database provides sea level data throughout all the domain, the statistical model presented here also finds

³⁹⁴ the relationship between the atmospheric forcing and the oceanic response at those areas.

All in all, we have proved here the good ability of the statistical model to reproduce daily maximum surges. The global obtained database provides daily

- ³⁹⁸ surge values for 140 years, and can be freely used by the scientific community to study, for instance, the long term-variability (decadal variations) during the
- ⁴⁰⁰ 20th Century, similarly to what it is done by Marcos et al. 2015 for tide gauge extremes. An interesting application would also be the use of these data as hy-
- ⁴⁰² drodynamic boundary conditions for numerical downscaling studies. An added value of this method is not only the database but also the tailor-made pre-
- dictors for each specific location worldwide; meaning that this study provides daily sea level pressure-induced predictors that can be used as covariates in
- ⁴⁰⁶ non-stationary or time-dependent models (Méndez et al. 2007; Menéndez and Woodworth 2010; Serafin and Ruggiero 2014), or that can also be used to set the
- 408 statistical model using surge results from regional models that would estimate maximum values with more accuracy.
- Future work regarding the improvement of the predictor's resolution will be carried out at specific areas to show how this affects the predicted surge. This
- ⁴¹² methodology could also be applied to regional climate projections of surge levels for different climate change scenarios.

414 Acknowledgements

This work has been partially funded by the Spanish Ministry of Economy and Competitiveness under the research project PLVMA (TRA2011- 28900). The authors would like to thank the University of Hawaii Sea Level Center for the tide gauge data. The DAC data (Dynamic Atmospheric Correction) are produced by CLS Space Oceanography Division and distributed by Aviso, with support from Cnes (http://www.aviso.altimetry.fr/). 20th Century Reanalysis V2 are data provided by the NOAA/OAR/ESRL PSD from their Web site at http://www.esrl.noaa.gov/psd/.

References

⁴²⁴ Calafat, F.M., Gomis, D., 2009. Reconstruction of Mediterranean sea level fields for the period 1945-2000. Global and Planetary Change 66, 225–234.

doi:10.1016/j.gloplacha.2008.12.015.

Camus, P., Méndez, F.J., Losada, I.J., Menéndez, M., Espejo, A., Pérez, J.,

428 Rueda, A., Guanche, Y., 2014a. A method for finding the optimal predictor indices for local wave climate conditions. Ocean Dynamics 64, 1025–1038.

Camus, P., Menéndez, M., Méndez, F.J., Izaguirre, C., Espejo, A., Cánovas, V.,

- ⁴³² Pérez, J., Rueda, A., Losada, I.J., Medina, R., 2014b. A weather-type statistical downscaling framework for ocean wave climate. Journal of Geophysical
- 434 Research: Oceans 119, 7389–7405. doi:10.1002/2014JC010141.

Carrère, L., Lyard, F., 2003. Modeling the barotropic response of the global
ocean to atmospheric wind and pressure forcing - comparisons with observations. Geophysical Research Letters 30, 1275. doi:10.1029/2002GL016473.

- ⁴³⁸ Casas-Prat, M., Wang, X.L., Sierra, J.P., 2014. A physical-based statistical method for modeling ocean wave heights. Ocean Modelling
- 440 73, 59-75. URL: http://www.sciencedirect.com/science/article/pii/ S1463500313001893, doi:10.1016/j.ocemod.2013.10.008.

⁴³⁰ doi:10.1007/s10236-014-0737-2.

- ⁴⁴² Compo, G.P., Whitaker, J.S., Sardeshmukh, P.D., Matsui, N., Allan, R.J., Yin, X., Gleason, B.E., Vose, R.S., Rutledge, G., Bessemoulin, P., BroNnimann,
- 444 S., Brunet, M., Crouthamel, R.I., Grant, a.N., Groisman, P.Y., Jones, P.D., Kruk, M.C., Kruger, a.C., Marshall, G.J., Maugeri, M., Mok, H.Y., Nordli,
- O., Ross, T.F., Trigo, R.M., Wang, X.L., Woodruff, S.D., Worley, S.J., 2011.
 The Twentieth Century Reanalysis Project. Quarterly Journal of the Royal
- 448 Meteorological Society 137, 1–28. doi:10.1002/qj.776.

Dangendorf, S., Müller-Navarra, S., Jensen, J., Schenk, F., Wahl, T., Weisse,

- R., 2014. North sea storminess from a novel storm surge record since AD
 1843. Journal of Climate 27, 3582–3595. doi:10.1175/JCLI-D-13-00427.1.
- ⁴⁵² Dee, D.P., Uppala, S.M., Simmons, a.J., Berrisford, P., Poli, P., Kobayashi, S., Andrae, U., Balmaseda, M.a., Balsamo, G., Bauer, P., Bechtold, P., Beljaars,
- a.C.M., van de Berg, L., Bidlot, J., Bormann, N., Delsol, C., Dragani, R.,
 Fuentes, M., Geer, a.J., Haimberger, L., Healy, S.B., Hersbach, H., Hólm,
- E.V., Isaksen, L., Kållberg, P., Köhler, M., Matricardi, M., Mcnally, a.P., Monge-Sanz, B.M., Morcrette, J.J., Park, B.K., Peubey, C., de Rosnay, P.,
- ⁴⁵⁸ Tavolato, C., Thépaut, J.N., Vitart, F., 2011. The ERA-Interim reanalysis: Configuration and performance of the data assimilation system. Quarterly
- Journal of the Royal Meteorological Society 137, 553–597. doi:10.1002/qj. 828.
- ⁴⁶² Espejo, A., Camus, P., Losada, I.J., Méndez, F.J., 2014. Spectral Ocean Wave Climate Variability Based on Atmospheric Circulation Patterns. Journal of
- 464 Physical Oceanography 44, 2139-2152. URL: http://journals.ametsoc. org/doi/abs/10.1175/JPO-D-13-0276.1, doi:10.1175/JPO-D-13-0276.1.
- 466 Giorgi, F., Christensen, J., Hulme, M., von Storch, H., Whetton, P., Jones, R., Mearns, L., Fu, C., Arritt, R., Bates, B., Benestad, R., Boer, G., Buishand,
- 468 A., Castro, M., Chen, D., Cramer, W., Crane, R., Crossly, J., Dehn, M., Dethloff, K., Dippner, J., Emori, S., Francisco, R., Fyfe, J., Gerstengarbe,
- 470 F., Gutowski, W., Gyalistras, D., Hanssen-Bauer, I., Hantel, M., Hassell, D.,

Heimann, D., Jack, C., Jacobeit, J., Kato, H., Katz, R., Kauker, F., Knutson, T., Lal, M., Landsea, C., Laprise, R., Leung, L., Lynch, A., May, W., Mc-472 Gregor, J., Miller, N., Murphy, J., Ribalaygua, J., Rinke, A., Rummukainen, M., Semazzi, F., Walsh, K., Werner, P., Widmann, M., Wilby, R., Wild, M., 474 Xue, Y., 2001. Regional Climate Information- Evaluation and Projections. URL: http://epic.awi.de/4973/. 476 Krueger, O., Schenk, F., Feser, F., Weisse, R., 2013. Inconsistencies between long-term trends in storminess derived from the 20CR reanaly-478 Journal of Climate 26, 868-874. URL: http: sis and observations. //journals.ametsoc.org/doi/abs/10.1175/JCLI-D-12-00309.1, doi:10. 480 1175/JCLI-D-12-00309.1. Lynch, D.R., Gray, W.G., 1979. A wave equation model for fi-482 element tidal computations. Computers & Fluids 7, 207nite

- 484 228. URL: http://www.sciencedirect.com/science/article/pii/ 0045793079900379, doi:10.1016/0045-7930(79)90037-9.
- 486 Marcos, M., Calafat, F.M., Berihuete, Á., Dangendorf, S., 2015. Long-term variations in global sea level extremes. Journal of Geophysical Research: Oceans
- 488 , n/a-n/aURL: http://doi.wiley.com/10.1002/2015JC011173, doi:10. 1002/2015JC011173.
- ⁴⁹⁰ Méndez, F.J., Menéndez, M., Luceño, A., Losada, I.J., 2007. Analyzing monthly extreme sea levels with a time-dependent GEV model. Journal of Atmospheric
- ⁴⁹² and Oceanic Technology 24, 894–911. doi:10.1175/JTECH2009.1.
- Menéndez, M., Woodworth, P.L., 2010. Changes in extreme high water levels
 based on a quasi-global tide-gauge data set. Journal of Geophysical Research:
 Oceans 115, 1–15. doi:10.1029/2009JC005997.
- ⁴⁹⁶ Müller-Navarra, S.H., Giese, H., 1999. Improvements of an empirical model to forecast wind surge in the German Bight. Deutsche Hydrographis-
- ⁴⁹⁸ che Zeitschrift 51, 385-405. URL: http://link.springer.com/10.1007/ BF02764162, doi:10.1007/BF02764162.

Pawlowicz, R., Beardsley, B., Lentz, S., 2002. Classical tidal harmonic analysis including error estimates in MATLAB using TDE. Computers and Geosciences 28, 929–937. doi:10.1016/S0098-3004(02)00013-4.

Serafin, K.A., Ruggiero, P., 2014. Simulating extreme total water levels us-

ing a time-dependent, extreme value approach. Journal of Geophysical Research: Oceans 119, 6305-6329. URL: http://doi.wiley.com/10.1002/
 2014JC010093, doi:10.1002/2014JC010093.

Wang, X.L., Feng, Y., Compo, G.P., Swail, V.R., Zwiers, F.W., Allan, R.J.,

 Sardeshmukh, P.D., 2013. Trends and low frequency variability of extratropical cyclone activity in the ensemble of twentieth century reanalysis. Cli mate Dynamics 40, 2775–2800. URL: http://link.springer.com/10.1007/

s00382-012-1450-9, doi:10.1007/s00382-012-1450-9.

- ⁵¹² Wang, X.L., Feng, Y., Swail, V.R., 2012. North Atlantic wave height trends as reconstructed from the 20th century reanalysis. Geophysical Research
- 514 Letters 39, n/a-n/a. URL: http://doi.wiley.com/10.1029/2012GL053381, doi:10.1029/2012GL053381.
- ⁵¹⁶ Wang, X.L., Swail, V.R., Cox, A., 2010. Dynamical versus statistical downscaling methods for ocean wave heights. International Journal of Clima-
- tology , n/a-n/aURL: http://doi.wiley.com/10.1002/joc.1899, doi:10. 1002/joc.1899.
- ⁵²⁰ Woodworth, P.L., Menéndez, M., 2015. Changes in the mesoscale variability and in extreme sea levels over two decades as observed by satellite altimetry.
- Journal of Geophysical Research: Oceans 120, 64-77. URL: http://doi. wiley.com/10.1002/2014JC010363, doi:10.1002/2014JC010363.

Highlights

- Novel combination of numerical and statistical modelling to reconstruct storm surges.
- Global reconstruction of storm surges for 140 years from 20th Century Reanlaysis (1871-2010).
- Global validation of the reconstructed storm surge with the non tidal residual from tide gauges.

A CER MAR